CENG 222 Statistical Methods for Computer Engineering

Week 6

Chapter 5 Computer Simulations and Monte Carlo Methods

Outline

- Generation of random numbers from specific distributions
 - Discrete distributions
 - Continuous distributions
- Chebyshev's inequality (3.3.7)
- Solving problems by Monte Carlo methods
 - Estimating probabilities
 - Estimating means and standard deviations

Uniform Random Numbers

- Tables of random numbers
- Pseudo-random number generators
 - Long sequences of random-looking numbers
 - Seed: starting location in the sequence
 - May use system time as seed
- Many systems provide standard uniform random number generators
 - Uniform(0,1)
- Question: Can we generate random numbers from any distribution using Uniform(0,1) rvs?

Bernoulli

• Let *U* be Uniform(0,1)

•
$$X = \begin{cases} 1, & \text{if } U$$

• P(success) = P(U < p) = p

Binomial

- Sum of *n* independent Bernoulli variables.
- Example
 - n = 20; p = 0.68;
 - U = rand(n, 1);
 - % generates an nx1 vector % of uniform random numbers X = sum(U < p);</pre>

Geometric

- Iterate and count the number of generated rvs until first success
- Example:

p = 0.16; X = 1; while rand > p; X = X+1; end;

Х

Negative Binomial

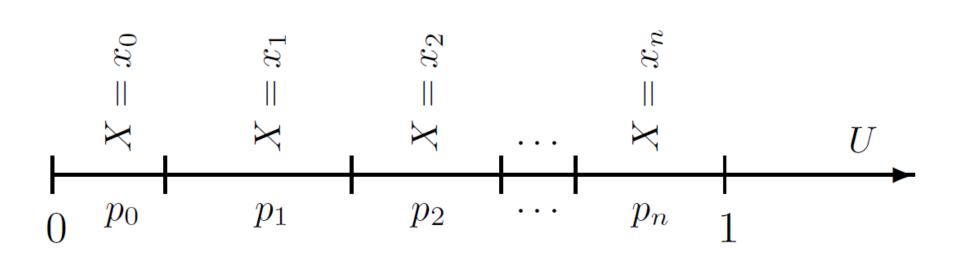
- Generate k independent Geometric(p) random numbers and sum them to get a NegativeBinomial(k,p) number.
- Example:

```
p = 0.16; X = 0; i = 0;
while i < k;
G = 1;
while rand > p;
G = G+1;
end;
X = X+G;
```

end;

• How efficient is generating a Binomial, a Geometric, or a Negative Binomial random number?

Arbitrary discrete distributions



Algorithm 5.1

1. Divide the interval [0,1] into subintervals A_i as follows:

 $- A_i = [p_0 + p_1 + ... + p_{i-1}, p_0 + p_1 + ... + p_i)$

- 2. Generate U, a standard uniform number
- 3. If *U* belongs to A_i then $X = x_i$
- How efficient is this method?
 - If you want to generate many Xs, efficiency is important.
 - O(*n*), O(log *n*), O(1)?
 - Check out the Alias Method, if you need an O(1) method.

Poisson

- Using Algorithm 5.1 to generate Poisson numbers.
- Example:

Inverse transform method

- Theorem: $U = F_X(X)$ is Uniform(0,1)
- Proof:
 - Note that the standard uniform cdf is $F_U(u) = u$ (i.e., $F'_U(u) = f_U(u) = 1$). We will try to show this fact using the given definition of $U = F_X(X)$

$$-F_U(u) = P(U \le u)$$

 $= P(F_X(X) \le u)$ = $P(X \le F_X^{-1}(u))$ = $F_X(F_X^{-1}(u))$

Inverse transform method

- If $U = F_X(X)$ then $X = F_X^{-1}(U)$
- The method:

- Generate a uniform random number

- Plug it in F_X^{-1} to generate X (i.e. solve for X).
- Example 5.10 (Exponential):

$$-F_{X}(X) = 1 - e^{-\lambda X} = U$$

$$- \Rightarrow X = -\frac{1}{\lambda} \ln(1 - U)$$

$$- \text{Can also use } X = -\frac{1}{\lambda} \ln(U) \text{ since } 1 - U \text{ is also}$$

$$\text{Uniform}(0, 1).$$

Inverse transform method

- Difficult to use if the inverse of the cdf is not easy to compute
- For example, for discrete distributions, $F_X^{-1}(U)$ does not exist. $U = F_X(X)$ has no roots, because X (hence $F_X(X)$) is finite and countable; whereas U is continuous.
- Therefore, for discrete rvs, we use the inverse method with a slight modification:
 - $-X = \min \{x \in S \text{ such that } F(x) > U\}$ where *S* is the set of possible values of *X*.

Example 5.12

• Using the inverse transform method for generating Geometric variables

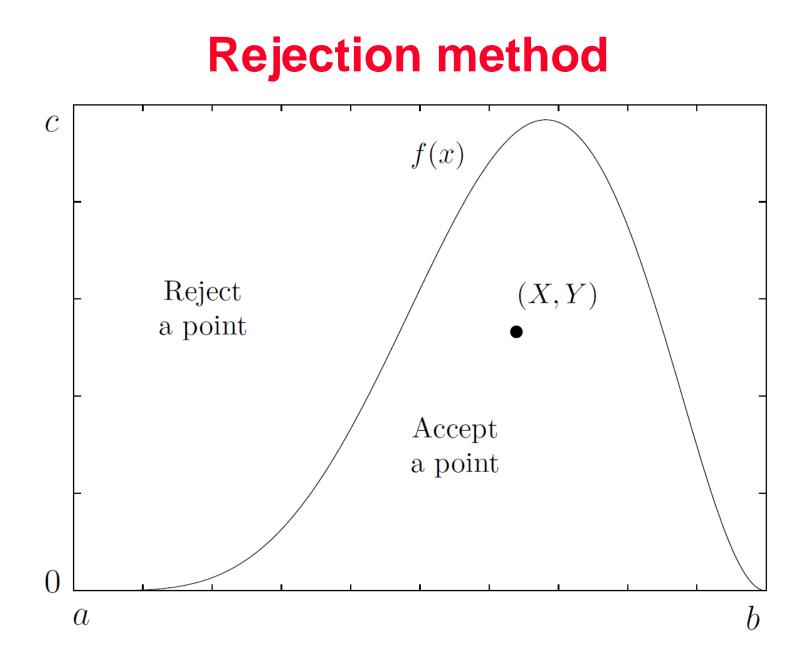
•
$$X = \left\lceil \frac{\ln(1-U)}{\ln(1-p)} \right\rceil$$

- The geometric variable is the ceiling of the exponential variable with $\lambda = -\ln(1-p)$
 - Exponential is the continuous analogue of geometric
 - Both have the memoryless property.

Rejection method

- When the cdf is difficult to solve for X and the pdf f_X is available, the rejection method can be used to generate random numbers from f_X .
- Idea:

- Generate 2D uniform coordinates (X,Y) in the bounding box of f_X and if $Y \leq f_X(X)$ output X.



Example

• The figure in the previous slide is the pdf of Beta(α =5.5, β =3.1)

$$-f_X = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \text{ for } 0 \le x \le 1$$

• Bounding box: m = 2.5, s = 0, t = 1.

a=5.5; b=3.1; s=0; t=1; m=2.5;

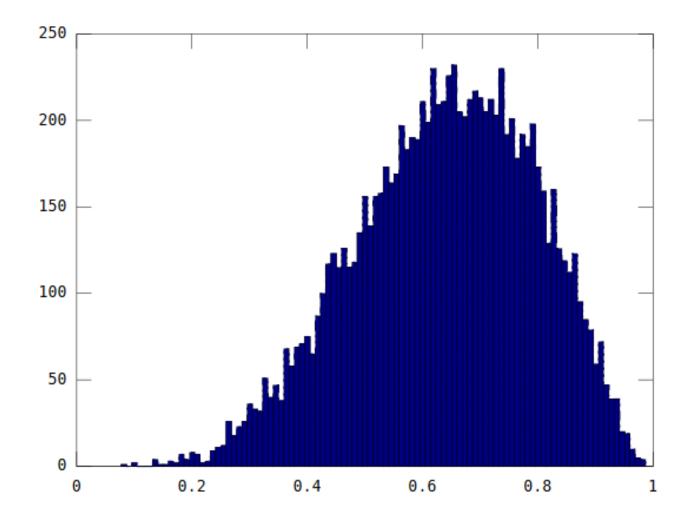
```
X = 0; Y = m;
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F = gamma(a+b)/gamma(a)/gamma(b)*X^(a-1)*(1-X)^(b-1); while (Y > F);

> U = rand; V = rand; X = s+(t-s)*U; Y = m*V; F = ... % same as above;

end; X

Example



Monte Carlo methods

- Generate many random variables from a distribution and estimate probabilities, means, standard deviations, etc. by simulating what happens in the long run.
- Question: How many numbers needed for acceptable results?
 - i.e., What will be the "size" of the Monte Carlo experiment?
 - Revisit Chebyshev's Inequality

Chebyshev's Inequality (3.3.7)

• For any distribution with expectation μ and variance σ^2 and for any positive ϵ

$$-P(|X-\mu| > \varepsilon) \le \left(\frac{\sigma}{\varepsilon}\right)^2$$

- In other words: any random variable *X* from the distribution is within ε distance of the μ with probability of at least $1 - (\sigma / \varepsilon)^2$

Estimating probabilities

- The probability *p*=*P*(*X* ∈ *A*) can be estimated as *p̂* by generating *N* random numbers and computing the proportion of random numbers that are in *A*.
- How accurate is the estimator?
 - What is $\mathbf{E}(\hat{p})$ and $\mathrm{Std}(\hat{p})$?
 - The number of X_i that are in *A* among the N generated random numbers is Binomial(*N*,*p*) with expectation *Np* and variance *Np*(1-*p*)
 - → $\mathbf{E}(\hat{p}) = p$ (unbiased estimator)

Std(\hat{p}) = $\sqrt{\frac{p(1-p)}{N}}$ the error in \hat{p} decreases with $1/\sqrt{N}$

How large should N be?

- Given the error ε and the probability, α , to exceed this error limit
- If an intelligent guess *p** on the value of *p* is available:

$$-N \ge p^*(1-p^*)\left(\frac{z_{\alpha/2}}{\varepsilon}\right)^2$$

• If not:

$$-N \ge 0.25 \left(\frac{z_{\alpha/2}}{\varepsilon}\right)^2$$

Example 5.14

How large should N be?

- If the *N* returned by these equations are not large enough for Binomial approximation, we may use Chebyshev's inequality:
 - If an intelligent guess p* on the value of p is available:

•
$$N \ge \frac{p^*(1-p^*)}{\alpha \varepsilon^2}$$

- If not:
• $N \ge \frac{1}{4\alpha \varepsilon^2}$

Estimating means and standard deviations

•
$$\overline{X} = \frac{1}{N} (X_1 + \dots + X_N)$$

– also unbiased and its error decreases with $1/\sqrt{N}$

•
$$s^{2} = \frac{1}{N-1} \sum_{i=1}^{N} (X_{i} - \bar{X})^{2}$$

- 1/N-1 needed so that $\mathbf{E}(s^{2}) = \sigma^{2}$